

ALMOST UNBIASED ESTIMATOR OF AN INEQUALITY MEASURE

APPENDIX

In order to derive the results given in Proposition 1 we first introduce, in matrix notation,

$$w = s^2 - \sigma^2 = \frac{1}{n-1} \hat{u}'\hat{u} - \sigma^2 = O_p\left(\frac{1}{\sqrt{n}}\right) \quad (\text{i})$$

$$\bar{u} = \frac{e'u}{n} = O_p\left(\frac{1}{\sqrt{n}}\right) \quad (\text{ii})$$

where u is an $n \times 1$ vector of u_i satisfying (2.2), e is an $n \times 1$ vector of unit elements and

$$M = I - \frac{ee'}{n} \quad (\text{iii})$$

is an $n \times n$ idempotent matrix with $\text{tr}(M) = n - 1$. Then from (2.1) and (2.4)

$$\hat{\theta} = \frac{s^2}{\bar{y}^2} = \frac{(w + \sigma^2)}{\mu^2} \left[1 + \frac{\bar{u}}{\mu}\right]^{-2} \quad (\text{iv})$$

Expanding the right hand side we get

$$\hat{\theta} - \theta = f_{-1/2} + f_{-1} + f_{-3/2} \quad (\text{v})$$

where, denoting f_{-r} as a term of $O_p(n^{-r})$

$$\begin{aligned} f_{-1/2} &= \frac{w}{\mu^2} - \frac{2\theta}{\mu} \bar{u} \\ f_{-1} &= -\frac{2}{\mu^3} w \bar{u} + \frac{3\theta}{\mu^2} \bar{u}^2 \\ f_{-3/2} &= \frac{3w\bar{u}^2}{\mu^4} - \frac{4\theta\bar{u}^3}{\mu^3} \end{aligned} \quad (\text{vi})$$

Thus, the bias of $\hat{\theta}$ to $O(n^{-1})$ is given,

$$\text{Bias}(\hat{\theta}) = E(f_{-1/2}) + E(f_{-1}) \quad (\text{vii})$$

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Now we state the following results:

$$\sigma^{-2} E[u'Bu] = trB \quad (\text{viii})$$

$$\sigma^{-3} E[u'Bu u] = \gamma_1(I * B)e \quad (\text{ix})$$

$$\sigma^{-4} E[u'Bu u u'] = \gamma_2(I * B) + (trB)I + 2B$$

$$\begin{aligned} \sigma^{-5} E\{u' Au u' Bu u\} &= \gamma_3(I * A * B)e + \gamma_1[(trA + 2A) \\ &\quad (I * B)e + (trB + 2B)(I * A)e + 4(I * AB)e] \end{aligned}$$

$$\begin{aligned} \sigma^{-6} E[u' Au . u' Bu . u u'] &= \gamma_4(I * A * B) + \gamma_2 [tr(A * B)I + tr A(I * B) \\ &\quad + trB(I * A) + 4(I * AB) + 2A(I * B) + \\ &\quad + 2B(I * A) + 2(I * B)A + 2(I * A)B] + \gamma_1^2[4(A * B) \\ &\quad + (I * A)ee'(I * B) + (I * B)ee'(I * A) \\ &\quad + 2I * \{A(I * B)\}ee' + 2I * \{B(I * A)\}ee'] \\ &\quad + 2(tr A)B + 4AB + 2(tr B)A + 4BA \\ &\quad + [2(tr AB) + tr(A) tr(B)]I. \end{aligned}$$

where A and B are n x n symmetric matrixes with non-stochastic elements—see, e.g. Ullah et. al. (1984, p. 398).

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Using (i), (ii), (viii), and (ix) it is easy to see that

$$\begin{aligned} Ef_{-1/2} &= 0 \\ Ef_{-1} &= \frac{1}{n}\theta^{3/2}(3\theta^{1/2} - 2\gamma_1) \end{aligned} \tag{xi}$$

which, when substituted in (vii) gives result (2.6) of the paper.

For obtaining the result (2.7), we write from (v), upto $O(n^{-2})$,

$$E(\hat{\theta} - \theta)^2 = E f_{-1/2}^2 + E f_{-1}^2 + 2E f_{-1/2} f_{-3/2} + 2E f_{-1/2} f_{-1}.$$

It is easy to verify that

$$\begin{aligned} E f_{-1/2}^2 &= \theta^2 \left(\frac{\gamma_2}{n} + \frac{2}{n-1} \right) - 4\theta^{5/2} \frac{\gamma_1}{n} + \frac{4\theta^3}{n} \\ E f_{-1}^2 + 2E f_{-1/2} f_{-3/2} &= \frac{\theta^3}{n^2} \left(10\gamma_2 \frac{n+3}{n-1} + \frac{20n}{n-1} - 96\theta^{1/2} \gamma_1 + 75\theta + 20\gamma_1^2 \frac{(n+1)}{(n-1)} \right) \\ 2E f_{-1/2} f_{-1} &= \frac{\theta^{5/2}}{n^2} \left(14\theta^{1/2} \gamma_2 \left(\frac{n}{n-1} \right) - 4\gamma_3 - 16\gamma_1 \frac{n}{n-1} - 12\theta \gamma_1 \right) \end{aligned}$$

To prove the result in proposition three, it is necessary to expand the sample skewness coefficient as well as the coefficient of variation squared. Letting $v = \hat{\mu} - \mu = \frac{1}{n} \sum \hat{u}_i^3 - \mu$ and

$$\hat{\gamma}_1 - \gamma_1 = g_{-1/2} + g_{-1} + g_{-3/2}$$

the mean squared error of $\tilde{\theta}$ is

$$\begin{aligned} MSE(\tilde{\theta}) &= MSE(\hat{\theta}) + \frac{1}{n^2} \{ 9\theta^4 + 4\theta^3 \gamma_1^2 - 12\gamma_1 \theta^{7/2} \} \\ &\quad - \frac{6}{n} \theta E \left(\theta f_{-1/2} + \theta f_{-1} + 2\theta f_{-1/2}^2 \right) - \frac{4}{n} \theta^{3/2} \gamma_1 E \left(f_{-1/2} + f_{-1} \right) \\ &\quad - \frac{6}{n} \theta^{1/2} \gamma_1 E f_{-1/2} - \frac{4}{n} \theta^{3/2} E f_{-1/2} g_{-1/2} + o \left(\frac{1}{n^2} \right) \end{aligned}$$

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and

$$Ef_{-1/2}g_{-1/2} = E \left[\frac{w}{\mu^2} - \frac{2\theta\bar{u}}{\mu} \right] \left[\frac{v}{\sigma^3} - \frac{3w\mu_3}{2\sigma^5} \right]$$

which upto $O(n^{-1})$, gives

$$Ef_{-1/2}g_{-1/2} = \frac{\theta}{n}\gamma_3 + \frac{6\theta}{n}\gamma_1 - \frac{2\theta^{\frac{3}{2}}}{n}\gamma_2 - \frac{3\theta}{n}\gamma_1\gamma_2 - \frac{3\theta}{n-1}\gamma_1 + \frac{3\theta^{\frac{3}{2}}}{n}\gamma_1^2.$$

Proposition 3 follows from straightforward algebra.