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Effects of renting on household energy expenditure: Evidence from Australia

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Abstract

This paper uses household survey data from Australia to investigate whether renters face larger energy bills than otherwise similar households. We find that a negative unconditional effect of renting on residential electricity expenditure becomes positive when controlling for log net wealth, with renters on average spending about 8% more than otherwise similar households. This is a larger effect than in most prior studies. The effect operates via higher usage quantities rather than higher average prices, and a similar effect is found for overall residential energy expenditure including natural gas. Central to the story is that renters tend to have lower net wealth, and net wealth is associated with higher energy use due to reasons including additional appliance ownership. This makes net wealth an important control. The findings cast light on the potential for more ambitious policy responses to reduce energy-related disadvantages faced by renters in Australia. There is also scope for further research into whether similarly large effects are evident in other countries.

Keywords:

energy expenditure; electricity consumption; wealth; rent; split incentive

JEL classification:

D10, G51, Q40, R20

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1. Introduction

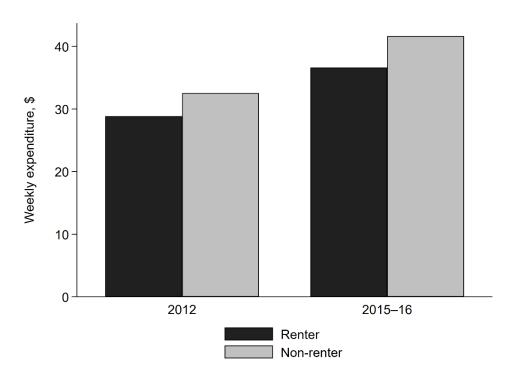
Renters typically do not have full information on the energy characteristics of a property that they may be considering, making them often unwilling to pay commensurately higher rent for more energy-efficient dwellings. As a result, landlords often have an inadequate incentive to invest in energy efficiency. There is a "split-incentive" or "principal-agent" problem. The result is that renters may face higher energy bills and/or difficulty controlling home temperatures, introducing a source of disadvantage for what is a relatively large share of the population in countries such as Australia (about 30%; ABS, 2017). Empirical analysis of split-incentive effects for renters has been identified as a promising area for research (Gerarden *et al.*, 2017).

This paper uses household survey data for Australia to explore whether renting households spend more on residential energy relative to otherwise similar households. We find that the additional residential energy expenditures of renters is larger than often realised. Our approach of including net wealth as a control could be utilized in follow-up studies to examine if a similarly sized effect exists in other countries.

Initial evidence, shown in Figure 1, indicates that as of 2012 renting households in Australia spent 11% *less* on grid-supplied electricity than non-renters (ABS, 2013). It can also be seen that this is similar for overall spending on electricity and domestic fuels including natural gas in 2015–16 (ABS, 2017). However, it is important to remember that renters differ from non-renters across various socio-economic dimensions, including average incomes and net wealth levels. This paper explores whether the negative unconditional effect of renting on household energy bills changes when key other factors are considered.

In prior research, some studies have found positive and significant effects of renting on energy use when including control variables, consistent with the existence of a split-incentive problem (Gerarden *et al.*, 2017). For example, Melvin (2018) found that renters' energy bills are approximately 2% higher due to split incentives in the United States (US). Best *et al.* (2021) found that renters in the US use approximately 9% more electricity than non-renters based on a model with many controls. Rehdanz (2007) found evidence of significantly higher expenditures on space heating and hot water supply for non-owner-occupiers in Germany, all else equal. Bernard *et al.* (1996) found that renters use more electricity in Quebec than homeowners, all else equal.

Figure 1. Average weekly household expenditure on electricity in 2012 and electricity and domestic fuel in 2015–16 for renters and non-renters.



Notes: The units are Australian dollars (A\$). Zeroes were excluded. The ABS calculates averages by dividing expenditure by the number of weeks in the recall period, which is generally quarterly based on this common billing frequency. Surveys were conducted throughout the year. Data: Australian Bureau of Statistics (ABS, 2013; 2017).

There is some contrasting evidence also. Meier and Rehdanz (2010) found a positive and significant coefficient for homeowners in the UK in explaining the log of annual heating expenditure per room, with some controls included. Studies for Australia have also found a negative effect of rental status on energy consumption (Wood *et al.*, 2012; Cao *et al.*, 2015; Poruschi and Ambrey, 2016) after conditioning on controls. Wood *et al.* (2012) concluded that the split-incentive problem may be less relevant in Australia due to landlords being adequately able to capture rent premia for energy-efficient properties. Australia maintains tax advantages (such as negative gearing allowances) for property investors, relatively unregulated rental markets without broad price controls, and a relatively high property turnover rate between owner-occupier and renter status (Wood *et al.*, 2012). Whether this context is sufficient to avoid split incentives is an open question.

Of the many other factors that may affect household consumption and expenditure on energy, net wealth may be particularly important. Higher net wealth means that households have a more relaxed budget constraint and the collateral to more easily borrow (Cooper and Dynan, 2016), and might lead to more energy consumption via three channels. First, wealthier

households may acquire a higher number of energy-using appliances. Second, the appliances they acquire may potentially use more energy, for example their televisions may be larger. Third, wealthier households may tend to use some appliances more intensively, for example their air conditioning system may be used more often or to cool a larger space.

There is a long history of studies that have documented wealth effects on various forms of consumption (Cooper and Dynan, 2016; Veirman and Dunstan, 2012), including in Australia (Windsor *et al.*, 2015; Atalay *et al.*, 2016). Both financial and nonfinancial wealth have been found to be positively associated with household consumption levels (Tan and Voss, 2003; Dvornak and Kohler, 2007). There is also a positive association between housing wealth and purchases of new passenger vehicles (Gillitzer and Wang, 2016), and wealth variables have been found to be relevant for investment in new energy generation capacity at various scales (Kim and Park, 2016; Best, 2017; Best *et al.*, 2019a). Yet to date there has been a lack of evidence on the effect of household wealth levels on household energy consumption. Prior studies on the effect of renting have yet to comprehensively control for household wealth effects.

A strong correlation between net wealth and tenure type is evident in Figure 2. In 2015–16, average net wealth was about A\$1.2 million for non-renting households and only about A\$0.25 million for renters. The correlation between net wealth and renting status was -0.2. Interestingly, the correlation between income and renting was weaker, at only -0.1.

Figure 3 shows that the energy burden (energy expenditure divided by income) is on average slightly higher for renters than non-renters. This is despite renters on average consuming less energy; according to the ABS (2013), renters on average consumed 15% less residential electricity than other households in 2012, for example. This points to a relative disadvantage: renters both consume less energy and have higher energy burdens than non-renters. Our econometric analysis controls for income, net wealth, and other factors to seek to identify whether renting places upward pressure on energy bills once potentially key variables are taken into account.

The conceptual framework for the paper is concise: while renters spend less on energy in an unconditional sense, this may be due to their lower average net wealth and incomes, among other factors. When accounting for these differences, renters may in fact spend more on energy relative to otherwise similar households, potentially reflecting the inferior energy

efficiency characteristics of rental properties and also perhaps other reasons. We investigate this issue using data from two separate household surveys.

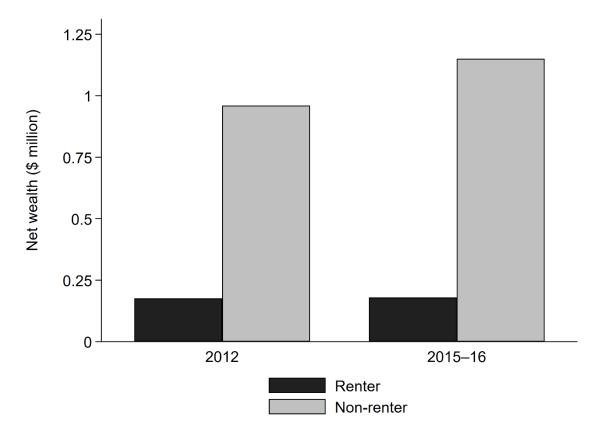


Figure 2. Household net wealth (A\$ million) for renters and non-renters.

Data: ABS (2013; 2017).

The analysis also explores the effects of many variables other than net wealth and income. Ownership of household appliances and equipment, which can be a positive function of net wealth, are likely to be relevant. For example, having rooftop solar photovoltaic panels helps households to reduce their use of grid-supplied electricity – and renters have been found to be substantially less likely to have solar panels, all else equal (Best *et al.*, 2019a, 2019b; Zander, 2020). In addition, there can be differences in behavioural measures between renters and non-renters (Nie *et al.*, 2020), so we include variables such as participation in a green power scheme. We also control for variables including whether a household is a single-parent household, the main source of household welfare such as age pensions, and whether there is a household member requiring energy-consuming medical equipment.

.03 - .02 - .01 -

Figure 3. Energy burden (energy expenditure divided by income).

Data: ABS (2013; 2017).

Our focus is partly motivated by the fact that renters are known to be more likely to experience energy-related financial stress in Australia (Best and Burke, 2019). Some already disadvantaged groups such as ethnic minorities and female-headed households are more likely to rent (ABS, 2017; Flage, 2018), heightening the importance of the issue. If there are potentials for improving energy efficiency in rental accommodation, this would deliver both socio-economic and environmental benefits. Evidence relating to a market failure in the rental market could thus carry policy relevance.

2. Method and data

2.1 Model and variables

The model in equation (1) considers household energy expenditure (E_h) as a function of renting, net wealth, and other variables:

$$\ln E_h = \alpha + \beta R_h + \gamma \ln W_h + X'_h \delta + \varepsilon_h \tag{1}$$

Separate regressions are run for residential electricity expenditure and for expenditure on residential energy including both electricity and natural gas.¹ Electricity bills include a supply

¹ Spending on energy for transport is not included.

charge that is independent of the quantity consumed plus a variable charge related to usage. In Australia, renters typically pay their own energy bills other than in a relatively small share of cases such as when a property is not separately metered. Indeed, about 94% of renters had non-zero electricity expenditure in the 2012 Household Energy Consumption survey that we use (ABS, 2013). Analysis in Section 3 uses the log of expenditure variables, so focuses on households with non-zero expenditures.

The 2012 Household Energy Consumption survey (ABS, 2013) provides data on electricity expenditure for all households and electricity usage for households who had their latest electricity bill readily available. As an additional dependent variable, we use the average electricity price paid by each household. We calculate this as the usage charge for grid-supplied electricity (excluding the fixed charge component) divided by the on-grid electricity usage quantity. We exclude the top 2% of responses (those above \$0.59 per kWh), as these appear to be inaccurate. An average price of \$0.59 per kWh is about three times the average and may in some cases reflect reporting errors. Electricity prices paid at the household level are partly affected by regional variation in average price levels, so we include location controls as described below.

We consider that net wealth and income are fundamental variables that may influence residential energy expenditure. There are two types of other variables. First are other household and housing characteristics that may be directly relevant for residential energy use, such as the number of people in the household. Second are variables that are potential direct channels via which net wealth and income may influence energy expenditure, including the types of hot-water systems used and the average hours of use of air conditioning. Like this study, the US study by Gillingham et al. (2012) controlled for characteristics of occupants, dwellings, and energy variables such as measures of heating system characteristics.

In equation (1), R_h is a binary variable equal to one for renters (defined at the household level) and zero otherwise. W_h is net wealth, equal to total assets minus total liabilities for each household. Assets include all financial, property, business assets, and other assets. Liabilities include loans for property, education, and credit card debt. Households with negative or zero values of net wealth are excluded when taking the log. This affects approximately 1% of households.

Other explanatory variables are in the X vector. These include a range of location, housing, socioeconomic, and energy variables. The location variables include binary variables for

regions. Climate zone controls are also included in some estimations. The housing variables include the tenure type, dwelling structure, and number of bedrooms in the residence to give an indication of size. Socioeconomic variables include household income along with a range of social variables such as single-parent status of households and also the age of the respondent who completed the survey on behalf of the household. Energy variables include a binary variable for having a mains gas connection, the principal substitute for electricity for residences. Table A.1 has a full list of the explanatory variables and their definitions. ε_h is the error term in what are ordinary least squares regressions. The results tables will show heteroskedasticity-robust standard errors.

2.2 Data

We use data from two household surveys conducted by the Australian Bureau of Statistics (ABS, 2013; 2017): the 2012 Household Energy Consumption (HEC) survey and the 2015–16 Household Expenditure Survey (HES). The advantage of these surveys is their wide coverage of energy variables. More recent surveys have included fewer energy-related questions. For example, the 2017–18 Survey of Income and Housing does not provide data on energy consumption. An alternative household survey, the Household Income and Labour Dynamics in Australia (HILDA) survey, does not include many energy variables (e.g. solar-panel uptake). Summary statistics for each variable are in Table A.2 and A.3.

2.3 Econometric and data issues

Sample selection bias could be an issue if the characteristics of the survey respondents are not representative of the Australian population. However, the two surveys were large and intended to be nationally representative. The HEC survey covered 11,978 households and the HES a sample of 10,046 households. We use the survey sampling weights in some regressions to assess the robustness of the main results. These weights are the inverse of the probability of a household being selected for the survey.

Variables such as net wealth and renting status tend to be relatively persistent over time, meaning that earlier data are likely to remain relevant in the current context. Similarity in the descriptive statistics for different surveys supports this expectation. For example, the proportion of renters is 30% in each of the two surveys used in this study, while the proportion of apartment dwellers is 9% in the earlier survey and 10% in the later survey. Dwellings on average had 3.09 bedrooms in the earlier survey and 3.07 bedrooms in the later survey. Both surveys use similar sampling approaches and were conducted by the same

statistical agency. The 2017–18 Survey of Income and Housing also has generally similar statistics.

Measurement error is a relevant consideration. For instance, net wealth could be subject to measurement error if respondents do not have accurate estimates of the values of their assets and liabilities. Taking the log of net wealth helps to remove a small number of outliers, such as households who report mortgages that are many times their home's value. As will be seen, similar coefficients are obtained using either survey.

Reverse causation is a potential problem for some variables. For example, electricity expenditure could influence appliance purchase decisions. This is unlikely to be a major issue though, as most energy-using or energy-generating appliances would have been installed in years prior to the survey year. We also conduct a number of robustness tests to reduce reverse causation concerns, including replacing a binary explanatory variable for households who have solar photovoltaic panels with a variable restricted to only panels installed at least two years prior to the survey. Causation or correlation between explanatory variables, such as renting and net wealth, is not a concern unless it is extremely high. Variance inflation factors will be examined to check for excessive multicollinearity.

The potential for omitted variable bias is another consideration. We attempt to reduce this by including many controls. These include measures of tenure, household financial resources and tendencies, household occupant characteristics, and building structure characteristics. Binary variables are also included for 14 regions: the Australian Capital Territory, the Northern Territory, and two for each of the six states using a capital city versus non-capital city split. We also control for the six climate zones available in the HEC data.

3. Results and discussion

3.1 2012 HEC survey

3.1.1 Main results

Column (1) of Table 1 uses the 2012 survey data, finding a negative coefficient for renting in explaining electricity expenditure.² This result is obtained while controlling for only location, climate, the survey timing over four quarters, and some socioeconomic and housing variables, including the number of people in the household. Column (2) adds a variable for

² Each column is for a constant sample to allow a focus on only the effects of including additional explanatory variables. This is achieved by excluding a small number of households with missing data for some variables from the estimation sample. Results available through the online code in the supplementary material are similar if this is not done.

the number of bedrooms, which tends to be higher for wealthier households. The coefficient for renting becomes less negative, moving from -0.116 to -0.068. Additional robustness tests in Appendix Table A.4 control for the number of electric appliances such as fridges. This also leads to the renter coefficient becoming less negative. However without the net wealth variable included as a control, the regressions still show a negative renter coefficient.

Columns (3)–(4) of Table 1 include additional energy controls. The renter coefficient becomes more negative in column (3) when having solar panels is included as a control, since renters are less likely to benefit from self-generation of solar electricity. Controlling for some other energy measures, such as the hours of use of air conditioning, leads to the renter coefficient becoming slightly less negative in column (4). Controlling for the number of TV set top boxes used at least once a week has minimal impacts, as evident via the Stata code.

Column (5) of Table 1 adds the log net wealth variable to the regression. Doing so results in the renter binary variable switching to being positive and significant. We refer to the coefficient here as the conditional direct effect of renting (Imai *et al.*, 2010). The result suggests that omitting net wealth in the earlier regressions led to the renter coefficient being biased downward because the renting variable was partly picking up the effect of lower net wealth among renters. The renter coefficient now has a magnitude of 0.075, suggesting that renters spend approximately 8% more on electricity than otherwise similar households who own their homes.

A positive coefficient is obtained for the net wealth variable in Table 1, suggesting that net wealth is relevant for residential electricity use in ways that have not been controlled for in the model. For example, wealthier households may tend to have larger residences that require greater electricity use for heating and cooling, and this is not fully controlled for by the variable measuring the number of bedrooms. Residence area in square meters has not been controlled for due to an absence of data for this variable.

The log net wealth coefficient in Table 1 is 0.07. This is an elasticity, as net wealth and electricity expenditure are both logged. The point estimate for the income elasticity of electricity expenditure is lower, at 0.04. Households who currently are more successful savers spend less on electricity, as indicated by negative and significant coefficients for those who just break even or usually save, relative to the excluded category of households who spend more than they receive. Robustness tests in the online code show similar coefficients for log

net wealth and the binary renter variable when the saving variables are excluded. Households with higher log housing costs also spend more on electricity.

There are many significant effects of occupant characteristics on electricity expenditure in Table 1. Households with more people spend more on electricity, as expected. The coefficient of 0.16 indicates that an additional person adds about 17% to average household electricity expenditure, all else equal.³ In column (5) there is a positive coefficient for single-parent households and a negative coefficient for households with a survey respondent aged 60 years or over.

The coefficients for the dwelling characteristics in Table 1 also match expectations. There is a positive effect of building age, as measured by a binary variable for houses aged 20 years or over, although statistical significance is mixed. Households in stand-alone houses spend more on electricity, while the effect for apartments is insignificantly different from that for the excluded category of medium-density townhouses in most columns. Having more bedrooms leads to higher electricity expenditure, as expected.

Column (5) of Table 1 indicates that having rooftop solar photovoltaic panels has a substantial impact on grid electricity expenditure net of feed-in tariffs, with a coefficient of – 0.53. This implies that households with solar panels on average spend approximately 41% less on their electricity purchases from the grid than households without solar panels, all else equal. We explored a number of robustness tests for this result, such as splitting the solar variable into two: a binary variable for solar panels in New South Wales (NSW) and the Australian Capital Territory (ACT) where gross rather than net feed-in tariffs were initially offered, and one for solar panels in other states. These give similar results. We also found similar results when the solar binary variable is restricted to solar installations from more than two years prior to the survey (to reduce possible reverse causation issues).

Energy-related explanatory variables in column (5) of Table 1 also have significant coefficients. Households with an occupant who uses energy-consuming medical equipment on average spend more on electricity, all else equal. This is as expected. There is a negative coefficient for households with mains gas connections, since this provides a substitute for electricity for space heating, water heating, and cooking. Multicollinearity does not appear to

 $^{^{3}}$ 100 * ($e^{0.16}$ - 1). 4 100 * ($e^{-0.53}$ - 1).

be a major concern in Table 1, as the average variance inflation factor, along with the values of the variance inflation factor for the renting and net wealth variables, are well below 10.

Table 1. Results for household electricity expenditure, HEC survey, 2012.

	Dependent (1)	variable: Log l (2)	household weel	kly electricity ((4)	expenditure (5)
Τ	-0.116***	-0.068***	-0.103***	-0.085***	0.075***
Tenure type: renter ^					
	(0.024)	(0.024)	(0.023)	(0.023)	(0.027)
Household income, log	0.062***	0.056***	0.058***	0.057***	0.042***
	(0.008)	(0.007)	(0.007)	(0.007)	(0.007)
Save: just break even ^	-0.058***	-0.052***	-0.056***	-0.054***	-0.057***
	(0.018)	(0.018)	(0.017)	(0.017)	(0.017)
Save most weeks ^	-0.125***	-0.116***	-0.117***	-0.115***	-0.133***
	(0.019)	(0.018)	(0.018)	(0.018)	(0.018)
Weekly housing costs, log	0.071***	0.063***	0.059***	0.064***	0.062***
	(0.009)	(0.009)	(0.009)	(0.008)	(0.008)
Green power scheme ^	0.062**	0.064**	0.081***	0.081***	0.075***
•	(0.028)	(0.028)	(0.027)	(0.027)	(0.027)
Number of persons #	0.178***	0.153***	0.160***	0.157***	0.160***
r	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)
Number of employed persons #	0.037***	0.031***	0.031***	0.034***	0.019**
rumber of employed persons "	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)
Number of children under 15 #	-0.076***	-0.067***	-0.070***	-0.068***	-0.069***
Number of emidren under 13 #	(0.012)	(0.012)	(0.011)	(0.011)	(0.011)
Single perent A	0.012)	0.012)	0.011)	0.011)	0.057***
Single parent ^					
G 1	(0.022)	(0.022)	(0.022)	(0.021)	(0.021)
Survey respondent 60+ years ^	-0.055***	-0.063***	-0.058***	-0.063***	-0.089***
	(0.018)	(0.018)	(0.017)	(0.017)	(0.017)
Mortgage ^	-0.088***	-0.065***	-0.064***	-0.072***	-0.019
	(0.024)	(0.024)	(0.023)	(0.022)	(0.022)
Dwelling age: 20+ years ^	0.015	0.036***	0.022*	0.018	0.021*
	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)
Flat/unit/apartment ^	-0.061**	-0.016	-0.013	-0.028	-0.024
	(0.026)	(0.026)	(0.026)	(0.025)	(0.025)
Separate house ^	0.156***	0.083***	0.096***	0.092***	0.084***
•	(0.020)	(0.020)	(0.020)	(0.019)	(0.019)
Number of bedrooms #	` /	0.109***	0.113***	0.117***	0.101***
		(0.008)	(0.008)	(0.008)	(0.008)
Solar electricity ^		(0.000)	-0.520***	-0.528***	-0.531***
Solar electricity			(0.032)	(0.032)	(0.032)
Medical equipment ^			(0.032)	0.141***	0.137***
wedicar equipment				(0.033)	(0.033)
Maine and A				-0.162***	
Mains gas ^					
NT.4 1/1 C1				(0.022)	(0.022)
Net wealth of household, log					0.072***
					(0.006)
Additional controls	Yes	Yes	Yes	Yes	Yes
Other energy characteristics	No	No	No	Yes	Yes
Observations	10,861	10,861	10,861	10,861	10,861
R^2	0.272	0.283	0.322	0.364	0.372

Notes. ***, **, * show statistical significance at 1, 5 and 10 per cent level respectively. Robust standard errors are in brackets below the coefficients. Coefficients for constants are not shown. "Additional controls" include categorical variables for calendar quarter of survey, state-capital city regions, and climate zones. "Other energy characteristics" includes the list of variables, as described in Table A.1. We also control for Tenure type = "Other". A detailed education variable with 11 levels for the household respondent is also used. The coefficients for this are nearly all insignificant, especially when controlling for log net wealth. A denotes binary variables, # denotes integer variables.

Extra robustness tests in Appendix A.4 include a model with additional explanatory variables for fridges/freezers, the number of hot water systems, having a pool or spa, and the taking of energy-saving actions. The coefficient for renting remains similar when including these. A number of further robustness tests are available through the online code in the Supplementary section, producing similar effects. Specifically, adding log net wealth without any other controls makes the renting coefficient positive and significant. A substantial change in the renting coefficient is also observed when net wealth is adjusted to exclude the value of the family home, although the renting coefficient does not become significant.

3.1.2 Decomposing the effect

Table 2 examines effects on three separate dependent variables: the log of electricity expenditure, the log average electricity price (excluding fixed charges), and the log electricity usage quantity. The coefficients in the log expenditure regression can be decomposed into the sum of the coefficients in the log price and log quantity regressions. For example, the renter coefficient of 0.098 in column (1) equals 0.005 plus 0.093. The small and insignificant coefficient for renters in explaining the average electricity price in column (2) and the large positive coefficient in explaining electricity use in column (3) suggest that renters have (conditionally) higher electricity expenditure primarily because of greater usage quantities. The magnitude of the impact on electricity use is similar to a finding for the US (Best *et al.*, 2021).

Electricity consumers in Australia can generally choose their preferred retail pricing arrangement based on a range of market options. For this and other reasons, there can be variation in the average electricity prices paid by individual households.⁵ However, many of the other variables in column (2) of Table 2 are insignificant. Location is an exception, with each of 13 location variables having positive and significant coefficients at the 1% level relative to the excluded zone of the ACT (see the online code in the Supplementary material section). This reflects the fact that the ACT has relatively low electricity prices. The largest coefficients are for the two South Australian regions (capital city and other), consistent with South Australia's relatively high retail electricity prices as of the year of the survey.⁶

⁵ The mean for the electricity consumption price in Table A.2 is \$0.21/kWh, with a standard deviation of \$0.05/kWh.

⁶ The average electricity consumption price in the 2012 HEC survey was highest for South Australia at \$0.24/kWh and lowest for the ACT at \$0.15/kWh.

Column (2) of Table 2 also finds a negative and significant effect on the log electricity price for households with a survey respondent who was aged 60 years or more. These households tend to pay approximately 3% less in terms of their average electricity consumption price than other households, all else equal. This likely largely reflects access to seniors' discounts (Victorian State Government, 2019; Australian Government, 2021). A negative and significant coefficient is also evident in Table A.5 when households with an average electricity consumption price in the top 5% (rather than the top 2%) are excluded from the sample.

Table 2 controls for a more concise education variable than was used in Table 1, equal to one for university-educated respondents and zero otherwise. The coefficient is negative and significant in explaining electricity spending and use but not price. That university-educated respondents have lower electricity use when controlling for other differences was also found for the US (Best *et al.*, 2021). Scale effects are also evident in the positive and significant coefficients for the number of people and the number of bedrooms. The estimate for the netwealth elasticity of electricity use, of about 0.09, is similar to the corresponding elasticity for electricity expenditure.

Appendix Table A.6 presents evidence on one channel for the influence of net wealth on energy consumption. Specifically, we find a positive relationship between log net wealth and the number of fridges/freezers owned by households. On average, a standard deviation increase in log net wealth is associated with a 0.1 unit increase in the number of fridges/freezers that a household has. There is also a positive net wealth effect on having a pool or spa, as expected. These are energy-using appliances.

Column (4) of Table 2 includes the log electricity price as an explanatory variable. The negative and significant coefficient of about -0.6 is a cross-sectional estimate of the price elasticity of electricity demand. This is within the range of estimates for Australia and other countries (Narayan and Smyth, 2005; Cao *et al.*, 2015; Poruschi and Ambrey, 2016; Burke and Abayasekara, 2018; Burke and Kurniawati, 2018). Most of the other coefficients remain similar with the inclusion of this electricity price variable.

Reverse causation from electricity usage to the average electricity price (excluding the fixed charge) is possible. However it may not be a major concern because most electricity customers retain their provider and electricity contract for multiple years (Australian Competition and Consumer Commission [ACCC], 2018). While block-tariff pricing exists in

South Australia, which means that average price would be a function of quantity, the block increments are typically not large.⁷ Block pricing has been uncommon in much of the rest of Australia.

Table 2. Results for electricity spending, price, and consumption, HEC survey, 2012.

	Ln spend	Ln price	Ln use	Ln use
	(1)	(2)	(3)	(4)
Tenure type: renter ^	0.098***	0.005	0.093***	0.096***
	(0.032)	(0.014)	(0.032)	(0.037)
University education ^	-0.049***	0.001	-0.050***	-0.027*
•	(0.014)	(0.006)	(0.015)	(0.016)
Household income, log	0.043***	-0.005	0.048***	0.041***
_	(0.009)	(0.004)	(0.009)	(0.009)
Save: just break even ^	-0.057***	-0.015	-0.042**	-0.057**
J	(0.021)	(0.009)	(0.022)	(0.024)
Save most weeks ^	-0.155***	-0.010	-0.146***	-0.143***
	(0.022)	(0.009)	(0.022)	(0.024)
Weekly housing costs, log	0.051***	0.003	0.048***	0.054***
, , , ,	(0.010)	(0.005)	(0.010)	(0.011)
Green power scheme ^	0.024	0.003	0.020	0.027
1	(0.032)	(0.012)	(0.034)	(0.038)
Number of persons #	0.199***	0.005	0.195***	0.188***
r	(0.011)	(0.005)	(0.011)	(0.013)
Number of employed persons #	-0.000	0.001	-0.001	0.003
r Jun r	(0.011)	(0.005)	(0.011)	(0.012)
Number of children under 15 #	-0.083***	-0.007	-0.076***	-0.069***
	(0.013)	(0.006)	(0.014)	(0.015)
Single parent ^	0.077**	0.007	0.071**	0.075**
28 F	(0.030)	(0.012)	(0.032)	(0.035)
Survey respondent 60+ years ^	-0.027	-0.031***	0.004	-0.014
zur eg respondent oo : yeurs	(0.019)	(0.008)	(0.020)	(0.022)
Mortgage ^	0.010	0.011	-0.001	-0.005
Wiorgage	(0.025)	(0.012)	(0.026)	(0.030)
Dwelling age: 20+ years ^	0.040***	-0.007	0.047***	0.010
2 weiling age. 20 Fyears	(0.014)	(0.006)	(0.014)	(0.016)
Flat/unit/apartment ^	0.024	0.015	0.009	0.023
rad and apartment	(0.033)	(0.014)	(0.035)	(0.035)
Separate house ^	0.079***	-0.000	0.080***	0.107***
beparate nouse	(0.025)	(0.011)	(0.026)	(0.027)
Number of bedrooms #	0.114***	0.001	0.113***	0.115***
rumber of bedrooms "	(0.010)	(0.004)	(0.010)	(0.012)
Net wealth of household, log	0.090***	0.004	0.086***	0.095***
Net wearm of nousehold, log	(0.008)	(0.003)	(0.008)	(0.010)
Electricity price, log	(0.000)	(0.003)	(0.000)	-0.595***
Electricity price, log				(0.053)
Additional controls	Yes	Yes	Yes	(0.033) Yes
Other energy characteristics	No	No	No	Yes
	No	No No	No No	
Probability weights Observations	7,409	7,409	7,409	Yes 7,409
R^2				*
π -	0.348	0.171	0.328	0.451

Notes. ***, **, * show statistical significance at 1, 5 and 10 per cent level respectively. Robust standard errors are in brackets below the coefficients. Coefficients for constants are not shown. "Additional controls" include categorical variables for calendar quarter of survey, state-capital city regions, and climate zones. "Other energy characteristics" includes the list of energy variables, as described in Table A.1. We also control for Tenure type = "Other". ^ denotes binary variables, # denotes integer variables.

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⁷ Based on energy deal comparisons on https://www.energymadeeasy.gov.au.

3.2 2015–16 HES

Table 3 uses 2015–16 data from the HES, with the dependent variable now being log expenditure on electricity plus other domestic fuels, principally natural gas. A similar pattern is evident, with a negative and significant effect for renters in column (1) changing to a positive and significant effect when log net wealth is included in column (4). Renters are again found to spend approximately 8% more on residential energy, all else equal, when the log net wealth variable is added to the regression. This is similar to the finding in Table 1 for electricity expenditure.

Key explanatory variables in Table 3 also produce reasonable results. The net wealth elasticity of expenditure of 0.07 is similar to the corresponding finding in Table 1. The point estimate for the net wealth coefficient is again higher than for the income coefficient. The average effect of having solar panels in Table 3 is lower than in Table 1. This is expected, as the dependent variable now includes expenditure on natural gas and other fuels. There is still a substantial negative association between having solar panels and energy bills, with significance at the 1% level.

The coefficients for the other control variables in Table 3 are also reasonable. There are positive and significant coefficients for household income, with statistical significance at the 1% level in each case. Energy expenditure is lower for households who have a general tendency to save, as indicated by the negative and significant coefficients for the "save most weeks" variable relative to the reference category of spending more than received. Households with higher housing costs also tend to spend more on energy. A scale effect is also evident, as there are positive and significant coefficients for the number of people in the household. Dwelling structure also matters, as there is a positive impact on energy expenditure for separate houses and negative coefficients for apartments. Larger dwellings are also associated with higher expenditure, all else equal, as suggested by the positive coefficients for the number of bedrooms.

Variables for the main source of social assistance also have significant associations with energy expenditure in Table 3. There are negative and significant coefficients for binary variables identifying age pensions and unemployment or student allowances as the main source of social assistance. These are in relation to a base category of households receiving no social assistance. This suggests that age pension and unemployment or student allowance recipients spend less on energy on average, all else equal. In contrast, other major welfare

categories tend to have energy expenditure that is not statistically different to the base category, as identified by insignificant coefficients that are available through the code in the Supplementary section.

Table 3. Results for expenditure on electricity and domestic fuel, HES, 2015–16.

	Dependent variable: Log expenditure on electricity and			
	domestic fuel			
	(1)	(2)	(3)	
Tenure type: renter ^	-0.135***	-0.085***	-0.084**	0.076**
	(0.026)	(0.027)	(0.035)	(0.038)
University education ^	-0.048***	-0.054***	-0.032**	-0.046***
	(0.013)	(0.013)	(0.016)	(0.016)
Household income, log	0.055***	0.050***	0.070***	0.050***
	(0.009)	(0.009)	(0.012)	(0.011)
Save: just break even ^	-0.026	-0.021	-0.015	-0.010
	(0.020)	(0.019)	(0.024)	(0.024)
Save most weeks ^	-0.111***	-0.105***	-0.110***	-0.119***
	(0.021)	(0.020)	(0.026)	(0.026)
Weekly housing costs, log	0.075***	0.066***	0.038***	0.041***
, , , ,	(0.010)	(0.010)	(0.013)	(0.013)
Age pension ^	-0.199***	-0.195***	-0.150***	-0.137***
1.86 bension	(0.023)	(0.023)	(0.028)	(0.027)
Unemployment and student	-0.123***	-0.116***	-0.111***	-0.082**
allowances ^	(0.031)	(0.030)	(0.034)	(0.035)
Number of persons #	0.199***	0.174***	0.167***	0.163***
rumber of persons "	(0.011)	(0.011)	(0.013)	(0.013)
Number of employed persons #	0.001	0.000	-0.001	-0.005
Number of employed persons #	(0.011)	(0.011)	(0.014)	(0.013)
Number of children under 15 #	-0.093***	-0.080***	-0.069***	-0.062***
Number of children under 13 #		(0.013)		
Cincle mannet A	(0.013)		(0.015)	(0.015)
Single parent ^	0.044*	0.031	0.016	0.036
3.6	(0.026)	(0.026)	(0.029)	(0.029)
Mortgage ^	-0.088***	-0.069***	-0.006	0.042
T1 . / /	(0.026)	(0.026)	(0.034)	(0.034)
Flat/unit/apartment ^	-0.103***	-0.061**	-0.079**	-0.069**
	(0.026)	(0.026)	(0.032)	(0.032)
Separate house ^	0.131***	0.065***	0.069***	0.062***
	(0.019)	(0.019)	(0.022)	(0.022)
Number of bedrooms #		0.098***	0.092***	0.078***
		(0.010)	(0.013)	(0.013)
Solar electricity ^			-0.437***	-0.439***
			(0.028)	(0.028)
Net wealth of household, log				0.069***
-				(0.009)
Additional controls	Yes	Yes	Yes	Yes
Probability weights	No	No	Yes	Yes
Observations	9,495	9,495	9,495	9,495
R^2	0.263	0.273	0.312	0.320

Notes. ***, **, * show statistical significance at 1, 5 and 10 per cent level respectively. Robust standard errors are in brackets below the coefficients. Coefficients for constants are not shown. Other social assistance categories are controlled for, with coefficients available through the Supplementary section; the reference category is "no social assistance". "Additional controls" include categorical variables for calendar quarter of survey and state-capital city regions. We also control for Tenure type = "Other". ^ denotes binary variables, # denotes integer variables.

3.3 Discussion

The positive net wealth effects found in this paper are consistent with the fundamental notion that consumption and investment will be greater when households have greater accumulated net wealth to finance their purchases. This is particularly relevant in the residential energy context given the large up-front costs of some energy-using appliances. Our results suggest that it can be difficult to accurately estimate the renter effect on residential energy expenditure without accounting for such net wealth effects.

Our main practical contribution is quantifying the renter effect in Australia. The estimated 8% higher energy bills for renters relative to otherwise similar households is large in magnitude – in dollar terms, it works out to be about A\$150 per year. This represents a specific additional burden for renters, some of whom are already under financial pressure. The paper feeds in to the understanding of factors contributing to residential energy insecurity, which is also a pressing concern in other countries like the US (Graff and Carley, 2020).

4. Conclusion and policy implications

While the residential energy expenditure of Australian renters is on average lower than for non-renters, this paper finds that a primary reason for this is that renters on average have lower net wealth. After controlling for log net wealth, we found positive effects of renting on residential energy expenditures. Specifically, renters spend approximately 8% more on electricity than otherwise similar households. The effect on overall residential energy expenditure (including natural gas) is similar. The effect on electricity bills is primarily due to higher electricity usage rather than higher average prices.

The apparent renter effect is likely to be at least in part due to a split incentive between renters and landlords resulting from information asymmetries. If this leads to rental properties being less energy efficient, it places upward pressure on the energy use of renters. There may also be other explanations, such as differences in socio-demographic factors and attitudes between renters and non-renters. However we have controlled for a range of socio-economic characteristics, a behavioural variable for air conditioning use in some estimations, and for the taking of energy-saving actions (see Appendix).

The apparently inferior energy efficiency of rental properties is a problem that can be difficult to address. Underinvestment by landlords is not only related to the famous principal-agent problem but also reflects broader issues such as a lack of landlord knowledge about energy

efficiency (Ambrose, 2015). There is also often a lack of engagement from tenants, especially when the size of potential energy savings is not well understood or when tenant tenure is not long (Palm, 2013). Government policy related to residential energy efficiency has also often not focused on landlords (Hope and Booth, 2014). Policymakers face the challenge of designing policies with sufficient scale and endurance, that allow for a fair distribution of incentives between tenants and landlords, and that have low transactions costs and sufficient political feasibility (Bird and Hernández, 2012).

One option is for incentives to invest in insulation, efficient energy-use equipment, or solar panels to be targeted at rental properties. For rooftop solar panels, the state of Victoria has indeed recently begun to specifically target subsidies to rental properties (Solar Victoria, 2021). Tax credits are another option (Charlier, 2015). There is also the potential for policymakers to better facilitate market solutions, such as third-party ownership approaches. For example, solar panel leasing could be increasingly suitable for budget-constrained renters (Rai and Sigrin, 2013).

Some of our other findings also carry policy relevance. Higher electricity expenditure and use among households with a household member who needs energy-consuming medical equipment highlights the case for this group to receive ongoing financial support. Single-parent households also tend to have higher electricity expenditure and usage, all else equal, pointing to a need for ongoing policy attention.

Future research could potentially use a similar approach of controlling for net wealth in exploring whether the large conditional effect of renting on residential energy expenditure estimated here for Australia exists in other countries. Future research could also potentially incorporate additional attitudinal variables or do more to break down the net wealth effect into its components. Additional research directly examining the energy efficiency performance of rental properties would also be of interest. Work into the efficacy of approaches to reduce the principal-agent problem also has the potential to be highly beneficial.

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⁸ At the federal level in Australia, support is currently available via the Essential Medical Equipment Payment.

References

- ABS, 2017. 6523.0 Household Income and Wealth, Australia, 2015-16.
- ABS, 2013. 4670.0 Household Energy Consumption Survey, Australia, 2012.
- Ambrose, A.R., 2015. Improving energy efficiency in private rented housing: Why don't landlords act? Indoor Built Environ. 24, 913–924. https://doi.org/10.1177/1420326X15598821
- Atalay, K., Whelan, S., Yates, J., 2016. House prices, wealth and consumption: New evidence from Australia and Canada. Rev. Income Wealth 62, 69–91. https://doi.org/10.1111/roiw.12145
- Australian Competition and Consumer Commission [ACCC], 2018. Restoring electricity affordability and Australia's competitive advantage.
- Australian Government, 2021. Commonwealth Seniors Health Card. https://www.dva.gov.au/health-and-treatment/veteran-healthcare-cards/commonwealth-seniors-health-card, accessed 11 March 2022.
- Bernard, J., Bolduc, D., Bélanger, D., 1996. Quebec residential electricity demand: A microeconometric approach. Can. J. Econ. 29, 92–113.
- Best, R., 2017. Switching towards coal or renewable energy? The effects of financial capital on energy transitions. Energy Econ. 63, 75–83. https://doi.org/10.1016/j.eneco.2017.01.019
- Best, R., Burke, P.J., 2019. Factors contributing to energy-related financial stress in Australia. Econ. Rec. 95, 462–479. https://doi.org/10.1111/1475-4932.12504
- Best, R., Burke, P.J., Nishitateno, S., 2021. Factors Affecting Renters' Electricity Use: More Than Split Incentives. Energy J. 42 (5). https://doi.org/10.5547/01956574.42.5.rbes
- Best, R., Burke, P.J., Nishitateno, S., 2019a. Understanding the determinants of rooftop solar installation: evidence from household surveys in Australia. Aust. J. Agric. Resour. Econ. 63, 922–939. https://doi.org/10.1111/1467-8489.12319
- Best, R., Burke, P.J., Nishitateno, S., 2019b. Evaluating the effectiveness of Australia's small-scale renewable energy scheme for rooftop solar. Energy Econ. 104475. https://doi.org/10.1016/j.eneco.2019.104475
- Bird, S., Hernández, D., 2012. Policy options for the split incentive: Increasing energy efficiency for low-income renters. Energy Policy 48, 506–514. https://doi.org/10.1016/j.enpol.2012.05.053
- Burke, P.J., Abayasekara, A., 2018. The price elasticity of electricity demand in the United States: A three-dimensional analysis. Energy J. 39, 87–102. https://doi.org/10.5547/01956574.39.2.pbur
- Burke, P.J., Kurniawati, S., 2018. Electricity subsidy reform in Indonesia: Demand-side effects on electricity use. Energy Policy 116, 410–421. https://doi.org/10.1016/j.enpol.2018.02.018
- Cao, K., Mathews, R., Wang, S., 2015. Modelling household energy consumption using ABS survey data. Econ. Pap. 34, 36–47. https://doi.org/10.1111/1759-3441.12093
- Charlier, D., 2015. Energy efficiency investments in the context of split incentives among French households. Energy Policy 87, 465–479. https://doi.org/10.1016/j.enpol.2015.09.005
- Cooper, D., Dynan, K., 2016. Wealth effects and macroeconomic dynamics. J. Econ. Surv. 30, 34–55. https://doi.org/10.1111/joes.12090
- Dvornak, N., Kohler, M., 2007. Housing wealth, stock market wealth and consumption: A panel analysis for Australia. Econ. Rec. 83, 117–130. https://doi.org/10.1111/j.1475-4932.2007.00388.x
- Flage, A., 2018. Ethnic and gender discrimination in the rental housing market: Evidence from a meta-analysis of correspondence tests, 2006–2017. J. Hous. Econ. 41, 251–273.

- https://doi.org/10.1016/j.jhe.2018.07.003
- Gerarden, T., Newell, R., Stavins, R., 2017. Assessing the energy-efficiency gap. J. Econ. Lit. 55, 1486–1525. https://doi.org/10.1257/jel.20161360
- Gillingham, K., Harding, M., Rapson, D., 2012. Split incentives in residential energy consumption. Energy J. 33, 2, 37–62. https://doi.org/10.5547/01956574.33.2.3
- Gillitzer, C., Wang, J.C., 2016. Housing wealth effects: Cross-sectional evidence from new vehicle registrations. Econ. Rec. 92, 30–51. https://doi.org/10.1111/1475-4932.12277
- Graff, M., Carley, S., 2020. COVID-19 assistance needs to target energy insecurity. Nat Energy 5, 352–354.
- Hope, A.J., Booth, A., 2014. Attitudes and behaviours of private sector landlords towards the energy efficiency of tenanted homes. Energy Policy 75, 369–378. https://doi.org/10.1016/j.enpol.2014.09.018
- Imai, K., Keele, L., Yamamoto, T., 2010. Identification, inference and sensitivity analysis for causal mediation effects. Stat. Sci. 25, 51–71. https://doi.org/10.1214/10-STS321
- Kim, J., Park, K., 2016. Financial development and deployment of renewable energy technologies. Energy Econ. 59, 238–250. https://doi.org/10.1016/j.eneco.2016.08.012
- Meier, H., Rehdanz, K., 2010. Determinants of residential space heating expenditures in Great Britain. Energy Econ. 32, 949–959. https://doi.org/10.1016/j.eneco.2009.11.008
- Melvin, J., 2018. The split incentives energy efficiency problem: Evidence of underinvestment by landlords. Energy Policy 115, 342–352. https://doi.org/10.1016/j.enpol.2017.11.069
- Narayan, P.K., Smyth, R., 2005. The residential demand for electricity in Australia: An application of the bounds testing approach to cointegration. Energy Policy 33, 467–474. https://doi.org/10.1016/j.enpol.2003.08.011
- Nie, H., Kemp, R., Xu, J.H., Vasseur, V., Fan, Y., 2020. Split incentive effects on the adoption of technical and behavioral energy-saving measures in the household sector in Western Europe. Energy Policy 140. https://doi.org/10.1016/j.enpol.2020.111424
- Palm, J., 2013. Energy Efficiency in Tenant-Owners' Residences: The Process of Going from Objective to Implementation. Hous. Stud. 28, 57–73. https://doi.org/10.1080/02673037.2013.729266
- Poruschi, L., Ambrey, C.L., 2016. On the confluence of city living, energy saving behaviours and direct residential energy consumption. Environ. Sci. Policy 66, 334–343. https://doi.org/10.1016/j.envsci.2016.07.003
- Rai, V., Sigrin, B., 2013. Diffusion of environmentally-friendly energy technologies: Buy versus lease differences in residential PV markets. Environ. Res. Lett. 8. https://doi.org/10.1088/1748-9326/8/1/014022
- Rehdanz, K., 2007. Determinants of residential space heating expenditures in Germany. Energy Econ. 29, 167–182. https://doi.org/10.1016/j.eneco.2009.11.008
- Solar Victoria 2021. Welcome to the Solar Homes Program. https://www.solar.vic.gov.au/Tan, A., Voss, G., 2003. Consumption and wealth in Australia. Econ. Rec. 79, 39–56. https://doi.org/10.1111/1475-4932.00077
- Veirman, E., Dunstan, A., 2012. Debt dynamics and the relationship between consumption and cyclical wealth changes. Econ. Rec. 88, 330–340. https://doi.org/10.1111/j.1475-4932.2012.00812.x
- Victorian State Government, 2019. Concessions and rebates. https://www.victorianenergysaver.vic.gov.au/get-help-with-your-bills/concessions-and-rebates, accessed 11 March 2022.
- Windsor, C., Jääsklä, J.P., Finlay, R., 2015. Housing wealth effects: Evidence from an Australian panel. Economica 82, 552–577. https://doi.org/10.1111/ecca.12137
- Wood, G., Ong, R., McMurray, C., 2012. Housing tenure, energy consumption and the split-

incentive issue in Australia. Int. J. Hous. Policy 12, 439–469. https://doi.org/10.1080/14616718.2012.730218

Zander, K.K., 2020. Unrealised opportunities for residential solar panels in Australia. Energy Policy 142, 111508. https://doi.org/10.1016/j.enpol.2020.111508

Appendix

 Table A.1. Variable descriptions

Variable	Description
Electricity expenditure	Household weekly expenditure on electricity, net of solar PV feed-in tariff
Electricity expenditure	revenues.
Electricity consumption	Current weekly household electricity use (kWh).
Electricity price	Electricity consumption price, calculated as the electricity usage charge
, , , , , , , , , , , , , , , , , , ,	divided by electricity consumption, excluding the top 2% of responses
	(those above \$0.59 per kWh).
Tenure type: renter	Tenure type of rent = 1 , otherwise = 0 .
University education	In Table 1, there are 11 categories for the highest education level of the
·	respondent. In Table 2 and 3, a more concise binary variable is used that
	equals 1 for respondents with a university education.
Quarter fixed effects	Quarter of year of interview.
Region fixed effects:	
NSW, outside capital	New South Wales excluding capital city = 1 , otherwise = 0 .
VIC, outside capital	Victoria excluding capital city = 1 , otherwise = 0 .
QLD, outside capital	Queensland excluding capital city = 1 , otherwise = 0 .
SA, outside capital	South Australia excluding capital city = 1 , otherwise = 0 .
WA, outside capital	Western Australia excluding capital city = 1 , otherwise = 0 .
TAS, outside capital	Tasmania excluding capital city = 1 , otherwise = 0 .
NSW, capital city	New South Wales and capital city = 1 , otherwise = 0 .
VIC, capital city	Victoria and capital city = 1 , otherwise = 0 .
QLD, capital city	Queensland and capital city = 1, otherwise = 0.
SA, capital city	South Australia and capital city = 1, otherwise = 0.
WA, capital city TAS, capital city	Western Australia and capital city = 1, otherwise = 0 .
Northern Territory	Tasmania and capital city = 1, otherwise = 0 . Northern Territory = 1, otherwise = 0 .
ACT	Australian Capital Territory = 1, otherwise = 0 .
Climate zone fixed effects	Climate zone categories: 1. Zone 1 (High humidity summer, warm winter);
Cilliate Zolle fixed circets	2. Zone 2 (Warm humid summer, mild winter); 3. Zone 3/4 (Hot dry
	summer, cool or warm winter); 4. Zone 5 (Warm temperate); 5. Zone 6
	(Mild temperate); 6. Zone 7/8 (Cool temperate/ Alpine).
Household income	Total previous financial year household income from all sources.
Saving behaviour	Management of household income over the last 12 months: spend more
6	money than receive, just break even, or able to save money most weeks.
	These are separate binary variables.
Housing cost	Weekly housing costs (no tenure adjustment) in Australian dollars.
_	Recurrent outlays by household members for shelter including items such
	as mortgage repayments, rent, and water rates.
Green power scheme	Whether connected to accredited Green Power electricity, binary variable.
Number of persons	Number of persons in the household.
Number of employed persons	Number of employed persons in the household.
Number aged under 15	Number of dependent children aged under 15 years in the household.
Single parent	Single parent with dependent children, binary variable.
Survey respondent 60+ years	Binary variable =1 if age of household reference person is 60 or more.
Mortgage	Owned with a mortgage = 1, otherwise = 0 .
Tenure type: Other	Tenure type: other (not owner or renter), binary variable.
Dwelling age: 20+ years	Estimated age of dwelling: 20 years old or more = 1, otherwise = 0.
Flat/unit/apartment	Flat/unit/apartment, binary variable. Separate (i.e. stand-alone) house, binary variable.
Separate house Number of bedrooms	Number of bedrooms.
Net wealth	Net wealth of household is assets minus liabilities, in dollars. The largest
Tiot would!	contributors to household assets are residential property and
	superannuation. Zero and negative net wealth values are effectively
	excluded when taking the log. There are 142 net wealth values of zero or
	less out of the total sample of 11,978 in the 2012 HEC survey.
Solar electricity	Value = 1 for households that have solar panels to generate electricity, 0
,	otherwise.
	Outer wise.

Medical equipment Whether anyone in the household uses energy-consuming medical

equipment, binary variable.

Mains gas

Other energy characteristics:

Insulation Whether dwelling has insulation. Category 1: has insulation, category 2:

does not have insulation, category 3: don't know.

Binary variable = 1 when household uses mains gas.

Replaced heater or cooler/air conditioner (AC) to a more efficient model, Replaced air conditioner

binary variable.

Replaced major white-good to a more efficient model, binary variable. Replaced white-good Heating hours

Average hours per day of reverse-cycle air-conditioning (heat-pump)

Whether the household has a continuous electric (storage or instantaneous)

systems used for heating during winter, categorical variable.

Cooling hours Average hours used per day for reverse-cycle air-conditioning (heat-pump)

systems used for cooling during summer, categorical variable.

Continuous electric hot water

systems.

hot water system in the dwelling. Off peak Whether household has any off-peak electric hot water systems, binary. Electric heat pump Whether household has any electric heat pump hot water systems, binary. Mains gas hot water Whether household has hot water systems that uses mains gas as its main

source of energy.

LPG hot water Whether household has any hot water systems that use LPG / bottled gas as

its main source of energy, binary.

Solar hot water Whether household has any hot water systems that use solar energy as its

main source of energy, binary.

Electric booster Whether household has electric booster for solar hot water system(s),

Gas booster Whether household has mains gas booster for solar hot water system(s),

binary.

Other 'quantity' variables

Number of fridges Number of fridges/freezers of all types.

Pool and/or spa Whether household has a swimming pool and/or spa, binary variable. Energy saving actions Number of actions from members of household to reduce energy use (min.

0, max. 8) out of: low flow shower head, energy efficient light bulbs, short showers, cold water for clothes washing, drying clothes on line, switching appliances off at wall, switching off chargers when not in use, draft proof

seals on doors and windows.

Number of hot water systems in dwelling. Number of hot water systems

Number of hot water system(s) that use electricity as its source of energy. Hot water using electricity

Notes. These variables are from the 2012 Household Energy Consumption (HEC) survey (ABS, 2013). Some variables are unavailable for the Household Expenditure Survey (HES) (ABS, 2017). The energy expenditure variable for the HES also includes spending on domestic fuel. The HES includes a variable on the main source of social assistance, dividing households into the following groups: no social assistance, age pension, disability and carer payments, family support payments, unemployment and student allowances, and other government pensions and allowances.

Table A.2. Descriptive statistics: HEC survey, 2012

Variable	Mean	Standard deviation
Household weekly expenditure on electricity (\$)	31.45	21.87
Household weekly electricity consumption (kWh)	126.75	115.86
Electricity consumption price (\$/kWh)	0.21	0.05
Tenure type: renter ^	0.30	0.46
University education ^	0.35	0.48
Quarter of interview #	2.44	1.10
New South Wales, excluding capital city ^	0.07	0.25
Victoria, excluding capital city ^	0.08	0.26
Queensland, excluding capital city ^	0.07	0.26
South Australia, excluding capital city ^	0.08	0.28
Western Australia, excluding capital city ^	0.08	0.27
Tasmania excluding, capital city ^	0.07	0.25
New South Wales, capital city ^	0.10	0.31
Victoria, capital city ^ Queensland, capital city ^	0.09 0.07	0.29 0.26
South Australia, capital city ^	0.07	0.28
Western Australia, capital city ^	0.07	0.26
Tasmania, capital city ^	0.04	0.20
Northern Territory ^	0.03	0.19
Australian Capital Territory ^	0.06	0.23
Climate zone #	4.16	1.40
Household income, annual (\$)	85,330	87,001
Save: just break even ^	0.46	0.50
Save most weeks ^	0.41	0.49
Weekly housing costs (\$)	260.38	281.03
Whether connected to accredited Green Power electricity ^	0.04	0.20
Number of persons in household #	2.44	1.29
Number of employed persons in the household #	1.22	1.02
Number of dependent children aged under 15 years #	0.49	0.90
Single parent with dependent children ^	0.07	0.25
Survey respondent is 60+ ^	0.32	0.47
Mortgage ^	0.36	0.48
Dwelling age 20+ years ^	0.66	0.48
Flat/unit/apartment ^	0.09	0.28
Separate house ^	0.81	0.39
Number of bedrooms #	3.09	0.89
Net wealth of household (\$)	730,944	1,272,454
Solar electricity ^	0.10	0.30
Insulation: yes ^	0.74	0.44
Insulation: don't know ^	0.14	0.35
Energy-consuming medical equipment ^	0.03	0.16
Mains gas ^	0.44	0.50
Number of fridges of all types #	1.74	0.88
Pool and/or spa^	0.10	0.30
Replaced heater or AC to a more efficient model ^	0.09	0.28
Replaced major white-good to a more efficient model ^	0.22	0.42
Average hours use of reverse-cycle AC: winter # (categorical)	7.41	2.94
Average hours use of reverse-cycle AC: summer # (categorical) Number of energy-saving actions #	6.91 5.11	3.25 1.74
Number of energy-saving actions # Number of hot water systems in dwelling #	1.02	0.19
Number of hot water systems in dwening # Number of hot water system(s) that use electricity #	0.51	0.19
Continuous electric hot water system \(^{\}\)	0.25	0.44
Off-peak electric hot water systems ^	0.23	0.42
Electric heat pump hot water systems ^	0.23	0.10
Hot water system that uses mains gas ^	0.34	0.10
Hot water system that uses mains gas Hot water systems that use LPG / bottled gas ^	0.04	0.20
Hot water systems that use solar energy ^	0.11	0.20
Electric booster for solar hot water system(s) ^	0.08	0.28
Mains gas booster for solar hot water system(s) ^	0.01	0.11
Notes. ^ denotes binary variables, # denotes integer/categorical variable		

Notes. ^ denotes binary variables, # denotes integer/categorical variables. There are 11,978 observations, although some energy questions were not answered by some households (e.g. electricity consumption, number of fridges). Electricity expenditure and consumption excludes 682 and 13 zero responses respectively. We have provided statistics for households with positive income and net wealth as our results take the log of these variables, but this has little effect because only approximately 100 households have negative values.

Table A.3. Descriptive statistics: HES, 2015–16

Variable	Mean	Standard deviation
Weekly expenditure on electricity and domestic fuel (\$)	40.17	27.07
Tenure type: renter ^	0.30	0.46
University education ^	0.40	0.49
Quarter of interview #	2.75	1.12
New South Wales, excluding capital city ^	0.05	0.21
Victoria, excluding capital city ^	0.04	0.19
Queensland, excluding capital city ^	0.04	0.20
South Australia, excluding capital city ^	0.03	0.16
Western Australia, excluding capital city ^	0.02	0.15
Tasmania excluding, capital city ^	0.02	0.14
New South Wales, capital city ^	0.18	0.38
Victoria, capital city ^	0.20	0.40
Queensland, capital city ^	0.10	0.31
South Australia, capital city ^	0.10	0.31
Western Australia, capital city ^	0.10	0.30
Tasmania, capital city ^	0.06	0.23
Northern Territory ^	0.03	0.18
Australian Capital Territory ^	0.03	0.17
Household income, annual (\$)	89,607	114,943
Save: just break even ^	0.44	0.50
Save most weeks ^	0.43	0.50
Weekly housing costs (\$)	256.16	266.22
Number of persons in household #	2.34	1.26
Number of employed persons in the household #	1.06	1.02
Number of dependent children aged under 15 years #	0.43	0.86
Single parent with dependent children ^	0.07	0.25
Mortgage ^	0.33	0.47
Flat/unit/apartment ^	0.10	0.30
Separate house ^	0.78	0.41
Number of bedrooms #	3.07	0.92
Solar electricity ^	0.17	0.38
Net wealth of household (\$)	868,669	2,003,284
Age pension ^	0.27	0.44
Disability and carer payment ^	0.08	0.28
Family support payments ^	0.12	0.32
Unemployment and student allowances ^	0.05	0.22
Other government pensions and allowances ^	0.09	0.29

Notes. ^ denotes binary variables, # denotes integer/categorical variables. There are 10,046 observations. Electricity expenditure excludes 263 zero responses. 108 values of net wealth, 60 values of income, and 173 for housing costs are zero or below. These are dropped from the regressions since we take the log and are also dropped from this table.

Table A.4. Results for household electricity expenditure, HEC survey, 2012.

Tuble 71.4. Results 1			household wee.		
Tenure type: renter ^	-0.128***	-0.074***	-0.107***	-0.057**	0.074**
Tenure type. Tenter	(0.027)	(0.027)	(0.027)	(0.026)	(0.030)
Household income log	0.059***	0.027)	0.055***	0.043***	0.029***
Household income, log	(0.009)	(0.009)	(0.009)	(0.043)	(0.029)
Save: just break even ^	-0.045**	-0.042**	-0.042**	-0.039**	-0.035*
Save. Just bleak even	(0.021)	(0.020)	(0.020)	(0.019)	(0.019)
Save most weeks ^	-0.122***	-0.115***	-0.113***	-0.103***	-0.115***
Save most weeks	(0.022)	(0.021)	(0.021)	(0.020)	(0.020)
Weekly housing costs, log	0.022)	0.073***	0.069***	0.068***	0.064***
Weekly housing costs, log	(0.011)	(0.011)	(0.010)	(0.010)	(0.004)
Green power scheme ^	0.071**	0.069**	0.086***	0.085***	0.087***
Green power scheme	(0.034)	(0.034)	(0.033)	(0.031)	(0.031)
Number of persons #	0.180***	0.054)	0.163***	0.145***	0.150***
rumber of persons #	(0.010)	(0.011)	(0.010)	(0.010)	(0.010)
Number of employed persons #	0.040***	0.033***	0.032***	0.024**	0.010)
rumber of employed persons #	(0.011)	(0.011)	(0.010)	(0.010)	(0.012)
Number of children under 15 #	-0.081***	-0.072***	-0.077***	-0.070***	-0.074***
Trumber of emitten under 15 #	(0.014)	(0.014)	(0.013)	(0.012)	(0.012)
Single parent ^	0.032	0.020	0.006	0.036	0.058**
Single parent	(0.026)	(0.025)	(0.025)	(0.024)	(0.024)
Survey respondent 60+ years ^	-0.038*	-0.048**	-0.043**	-0.068***	-0.086***
Survey respondent 60+ years	(0.020)	(0.020)	(0.019)	(0.018)	(0.018)
Mortgage ^	-0.097***	-0.072***	-0.069***	-0.071***	-0.020
Mortgage	(0.028)	(0.027)	(0.026)	(0.025)	(0.025)
Dwelling age: 20+ years ^	0.018	0.040***	0.025*	0.019	0.021
Dwening age. 201 years	(0.014)	(0.014)	(0.014)	(0.013)	(0.013)
Flat/unit/apartment ^	-0.026	0.020	0.023	0.001	-0.004
Thay affic apartment	(0.029)	(0.029)	(0.029)	(0.028)	(0.028)
Separate house ^	0.162***	0.085***	0.101***	0.070***	0.064***
Separate nouse	(0.023)	(0.024)	(0.023)	(0.022)	(0.021)
Number of bedrooms #	(0.023)	0.117***	0.120***	0.091***	0.078***
Tighteen of economis		(0.010)	(0.009)	(0.009)	(0.009)
Solar electricity ^		(0.00-0)	-0.520***	-0.556***	-0.553***
zona ciccurcity			(0.035)	(0.034)	(0.034)
Medical equipment ^			(0.000)	0.122***	0.126***
1 1				(0.036)	(0.036)
Mains gas ^				-0.114***	-0.126***
Trans Bus				(0.025)	(0.025)
Number of fridges/freezers #				0.099***	0.094***
Trained of inages, needers				(0.007)	(0.007)
Pool and/or spa^				0.275***	0.262***
1 cor and or spa				(0.019)	(0.019)
Energy-saving actions #				(0.01)	-0.028***
Zheigy saving actions "					(0.004)
Net wealth of household, log					0.061***
1.00 weather of nousehold, log					(0.007)
Additional controls	Yes	Yes	Yes	Yes	Yes
Other energy characteristics	No	No	No	Yes	Yes
Observations	8,348	8,348	8,348	8,348	8,348
R^2	0.267	0.280	0.321	0.402	0.413
Notes *** ** * show statistical					andard errors are

Notes. ***, **, * show statistical significance at 1, 5 and 10 per cent level respectively. Robust standard errors are in brackets below the coefficients. Coefficients for constants are not shown. "Additional controls" include categorical variables for calendar quarter of survey, state-capital city regions, and climate zones. "Other energy characteristics" include the full list of energy variables, as described in Table A.1. We also control for Tenure type = "Other". There is also a detailed education variable with 11 levels for the household respondent as in Table 1; the coefficients are nearly all insignificant, especially when controlling for log net wealth. A denotes binary variables, # denotes integer variables.

Table A.5. Results for electricity price, HEC survey 2012, excluding households with the top 5% of average electricity price.

	Ln price
	(1)
Tenure type: renter ^	0.004
• •	(0.013)
University education ^	0.004
•	(0.006)
Household income, log	-0.000
	(0.003)
Save: just break even ^	-0.012
	(0.008)
Save most weeks ^	-0.011
	(0.008)
Weekly housing costs, log	0.004
	(0.004)
Green power scheme ^	0.005
	(0.012)
Number of persons #	0.002
X 1 6 1 1 "	(0.004)
Number of employed persons #	0.006
N 1 6 1 11 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	(0.004)
Number of children under 15 #	-0.002
Cin als mannet A	(0.005)
Single parent ^	0.009
Survey respondent 60+ years ^	(0.011) -0.017**
Survey respondent 60+ years	(0.008)
Mortgage ^	0.005
Mongage	(0.010)
Dwelling age: 20+ years ^	-0.005
D welling age. 20+ years	(0.005)
Flat/unit/apartment ^	0.014
Tius umu up urumom	(0.013)
Separate house ^	-0.003
	(0.010)
Number of bedrooms #	-0.000
	(0.004)
Net wealth of household, log	0.003
,	(0.003)
Additional controls	Yes
Observations	7,184
R^2	0.181

Notes. See Table 2.

Table A.6. Results explaining (1) the number of fridges/freezers and (2) a binary variable for having a pool and/or spa, 2012 HEC

	Fridge/freezer	Pool/spa
	(1)	(2)
Tenure type: renter ^	0.031	-0.020
	(0.041)	(0.013)
University education ^	-0.161***	-0.001
	(0.019)	(0.006)
Household income, log	0.037***	0.007*
	(0.010)	(0.004)
Save: just break even ^	0.035	-0.019**
	(0.028)	(0.009)
Save most weeks ^	-0.011	-0.024**
	(0.029)	(0.010)
Weekly housing costs, log	-0.020	0.011**
	(0.013)	(0.005)
Green power scheme ^	-0.056	0.002
	(0.039)	(0.014)
Number of persons #	0.156***	0.006
	(0.015)	(0.005)
Number of employed persons #	0.018	0.012**
	(0.015)	(0.005)
Number of children under 15 #	-0.130***	0.001
	(0.019)	(0.006)
Single parent ^	-0.005	-0.019*
	(0.034)	(0.010)
Survey respondent 60+ years ^	0.112***	0.004
	(0.027)	(0.008)
Mortgage ^	0.053	0.009
	(0.036)	(0.012)
Dwelling age: 20+ years ^	0.036*	0.013**
	(0.019)	(0.006)
Flat/unit/apartment ^	0.002	0.010
-	(0.032)	(0.008)
Separate house ^	0.189***	0.022***
•	(0.028)	(0.007)
Number of bedrooms #	0.098***	0.044***
	(0.014)	(0.004)
Net wealth of household, log	0.086***	0.021***
, 2	(0.009)	(0.003)
Additional controls	Yes	Yes
Observations	8,829	11,473
R^2	0.167	0.092

Notes. ***, **, * show statistical significance at 1, 5 and 10 per cent level respectively. Robust standard errors are in brackets below the coefficients. Coefficients for constants are not shown. "Additional controls" include categorical variables for calendar quarter of survey, state-capital city regions, and climate zones. We also control for Tenure type = "Other". ^ denotes binary variables, # denotes integer variables.